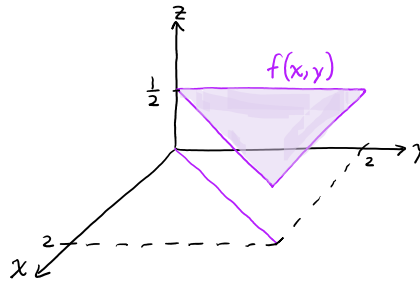
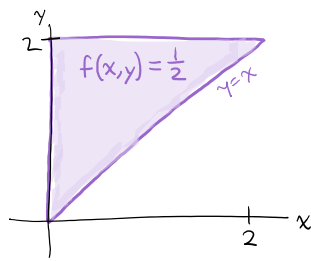
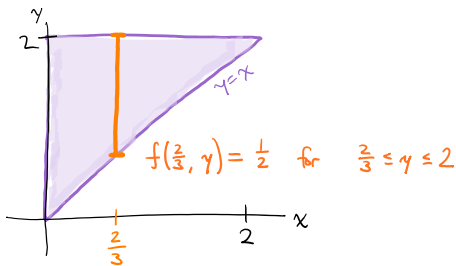


1. Let X and Y have joint density $f(x, y) = \frac{1}{2}$ for $0 \leq x \leq y \leq 2$.

(a) Sketch the joint density of X and Y .



(b) Suppose you know that $X = \frac{2}{3}$. Highlight $x = \frac{2}{3}$ on your sketch in part (a). How can you infer the density of Y given that $X = \frac{2}{3}$?



Since $f(\frac{2}{3}, y)$ is constant, the density of Y given $X = \frac{2}{3}$ is constant for $\frac{2}{3} \leq y \leq 2$. That is, if $X = \frac{2}{3}$ then Y is uniform on $[\frac{2}{3}, 2]$, so $f_{Y|X}(y | x = \frac{2}{3}) = \frac{3}{4}$ for $\frac{2}{3} \leq y \leq 2$.

(c) Let $0 \leq x_0 \leq 2$ and suppose you know that $X = x_0$. How can you infer the density of Y from $f(x_0, y)$?

Since $f(x_0, y)$ is constant for $x = x_0$, $x_0 \leq y \leq 2$, Y is uniformly distributed on $[x_0, 2]$.

Thus $f_{Y|X}(x_0, y) = \frac{1}{2 - x_0}$ for $x_0 \leq y \leq 2$.

(d) Find the marginal density $f_X(x)$. Use this to compute the conditional density $f_{Y|X}(y | x_0)$. Does this agree with your answer in part (c)?

$$\text{marginal density: } f_X(x) = \int_x^2 f(x, y) dy = \int_x^2 \frac{1}{2} dy = \frac{y}{2} \Big|_{y=x}^{y=2} = \frac{2-x}{2}$$

$$\text{then: } f_{Y|X}(y | x_0) = \frac{f(x_0, y)}{f_X(x_0)} = \frac{\frac{1}{2}}{\frac{2-x_0}{2}} = \frac{1}{2-x_0} \quad \text{for } x_0 \leq y \leq 2$$

and this agrees with our answer in part (c).

(e) What is the expected value of Y given that $X = x_0$?

$$E(Y | X = x_0) = \int_{x_0}^2 y \cdot \underbrace{f_{Y|X}(y|x_0)}_{\text{density}} dy = \int_{x_0}^2 y \cdot \frac{1}{2-x_0} dy = \frac{y^2}{2(2-x_0)} \Big|_{y=x_0}^{y=2}$$

$$= \frac{4-x_0^2}{2(2-x_0)} = \frac{2+x_0}{2}$$

↑ This is the mean of the Unif $[x_0, 2]$ distribution.

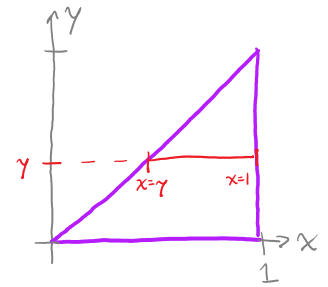
2. The joint pdf of X and Y is $f(x, y) = 3x$, for $0 \leq y \leq x \leq 1$.

(a) What is the conditional distribution of X given $Y = y$?

$$f_Y(y) = \int_y^1 3x dx = \frac{3}{2} x^2 \Big|_{x=y}^{x=1} = \frac{3}{2}(1-y^2) \quad \text{for } 0 \leq y \leq 1$$

$$f_{X|Y}(x|y) = \frac{f(x, y)}{f_Y(y)} = \frac{3x}{\frac{3}{2}(1-y^2)} = \frac{2x}{1-y^2} \quad 0 \leq y \leq x \leq 1$$

↑ fixed ↑ variable



(b) What is $E(X | Y = y)$?

$$E(X | Y = y) = \int_y^1 x \cdot \frac{2x}{1-y^2} dx = \frac{2}{1-y^2} \int_y^1 x^2 dx = \frac{2}{1-y^2} \cdot \frac{x^3}{3} \Big|_{x=y}^{x=1} = \frac{2}{1-y^2} \left(\frac{1}{3} - \frac{y^3}{3} \right) = \frac{2(1-y^3)}{3(1-y^2)}$$

(c) What is $\text{Var}(X | Y = y)$?

$$E(X^2 | Y = y) = \int_y^1 x^2 \cdot \frac{2x}{1-y^2} dx = \frac{1}{1-y^2} \cdot \frac{x^4}{2} \Big|_{x=y}^{x=1} = \frac{1-y^4}{2(1-y^2)} = \frac{1}{2}(1+y^2)$$

$$\text{Var}(X^2 | Y = y) = \frac{1}{2}(1+y^2) - \left(\frac{2(1-y^3)}{3(1-y^2)} \right)^2 = \frac{(y-1)^2 (y^2 + 4y + 1)}{18(1+y)^2}$$

3. For continuous random variables X and Y , show that $E(E(X|Y)) = E(X)$.

inside: conditional expectation
outside: expectation of a function of Y

$$E(X|Y=y) = \int_{-\infty}^{\infty} x \cdot f_{X|Y}(x|y) dx = \int_{-\infty}^{\infty} x \cdot \frac{f(x,y)}{f_Y(y)} dx, \text{ which is a function of } y$$

$$\begin{aligned} E(E(X|Y)) &= \int_{-\infty}^{\infty} E(X|Y=y) f_Y(y) dy = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} x \cdot \frac{f(x,y)}{f_Y(y)} dx f_Y(y) dy \\ &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} x f(x,y) dy dx = \int_{-\infty}^{\infty} x \left[\int_{-\infty}^{\infty} f(x,y) dy \right] dx \\ &= \int_{-\infty}^{\infty} x \cdot f_X(x) dx = E(X) \end{aligned}$$

4. The number of eggs N found in a nest of a certain species of turtle has a Poisson distribution with mean λ . Each egg has a probability p of being viable, and this event is independent from egg to egg. Find the mean and variance of the number of viable eggs per nest.

r.v.s: $N \sim \text{Poisson}(\lambda)$, $X = \text{number of viable eggs} \sim \text{Bin}(N, p)$

$$\text{mean: } E(X) = E(E(X|N)) = E(Np) = p E(N) = p \lambda$$

$E(X|N) = Np$ ↗

$$\begin{aligned} \text{variance: } \text{Var}(X) &= \text{Var}(E(X|N)) + E(\text{Var}(X|N)) = \text{Var}(Np) + E(Np(1-p)) \\ &= p^2 \text{Var}(N) + p(1-p)E(N) = p^2 \lambda + p(1-p) \lambda = p^2 \lambda - p^2 \lambda + p \lambda = p \lambda \end{aligned}$$

5. Simulate 10,000 averages, each of k samples from a Unif[0,1] distribution. Make a histogram of the 10,000 averages. Start with $k = 1$ and then try larger values of k . How does the shape of the histogram depend on k ?

We will do this next time.

6. Repeat the previous simulation, but now replace Unif[0,1] with a different distribution of your choice. What is the shape of the histogram? How does it depend on k ?

Next time...

BONUS: If X and Y are independent binomial random variables with identical parameters n and p , calculate the conditional expected value of X given that $X + Y = m$.

First, compute the conditional pmf of X given that $X + Y = m$.

$$\begin{aligned} P(X=k | X+Y=m) &= \frac{P(X=k \text{ and } X+Y=m)}{P(X+Y=m)} = \frac{P(X=k)P(Y=m-k)}{P(X+Y=m)} \\ &= \frac{\binom{n}{k} p^k (1-p)^{n-k} \cdot \binom{n}{m-k} p^{m-k} (1-p)^{n-m+k}}{\binom{2n}{m} p^m (1-p)^{2n-m}} \quad \left[\text{For the denominator, note that } X+Y \sim \text{Bin}(2n, p) \right] \\ &= \frac{\binom{n}{k} \binom{n}{m-k}}{\binom{2n}{m}} \end{aligned}$$

This is a hypergeometric probability: the probability of k successes in a sample of size m from a population with n successes and n failures.

So the conditional distribution of X , given that $X+Y=m$, is hypergeometric, and its mean is $E(X | X+Y=m) = \frac{m}{2}$.